

Aufgabe	A33	A34	A35	Σ
Punkte				

Aufgabe 33. (a) Anwenden der Formel der VL ergibt sofort für $x \in \mathbb{R}$:

$$\begin{aligned}
 f^X(x) &= \int_{\mathbb{R}} f^{X,Y}(x,y) \, dy \\
 &= \int_{\mathbb{R}} \frac{1}{\pi} \mathbb{1}_{\{(x,y) \in E\}} \, dy \\
 &= \mathbb{1}_{\{|x| \leq 1\}} \int_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}} \frac{1}{\pi} \, dy \\
 &= \frac{2}{\pi} \sqrt{1-x^2} \mathbb{1}_{\{|x| \leq 1\}}
 \end{aligned}$$

Ganz analog folgt für $y \in \mathbb{R}$:

$$f^Y(y) = \frac{2}{\pi} \sqrt{1-y^2} \mathbb{1}_{\{|y| \leq 1\}}.$$

(b) Anwenden der Formel der VL ergibt zunächst mit Anwendung des Transformationssatzes

$$\begin{aligned}
 \mathbb{E}(X) &= \int_{\mathbb{R}} x f^X(x) \, dx \\
 &= \frac{1}{\pi} \int_{-1}^1 2x \sqrt{1-x^2} \, dx \\
 &= \frac{1}{\pi} \left[\int_{-1}^0 2x \sqrt{1-x^2} \, dx + \int_0^1 2x \sqrt{1-x^2} \, dx \right] \\
 &\stackrel{z=1-x^2}{=} \frac{1}{\pi} \left[\int_0^1 -\sqrt{z} \, dz + \int_1^0 -\sqrt{z} \, dz \right] \\
 &= 0
 \end{aligned}$$

Unter erneuter Formelanwendung folgt

$$\begin{aligned}
 \mathbb{E}(X^2) &= \int_{\mathbb{R}} x^2 f^X(x) \, dx \\
 &= \frac{2}{\pi} \int_{-1}^1 x^2 \sqrt{1-x^2} \, dx \\
 &\stackrel{x=\sin(\varphi)}{=} \frac{2}{\pi} \int_{-\frac{\pi}{2}}^{\frac{\pi}{2}} \sin^2(\varphi) \cos^2(\varphi) \, d\varphi \\
 &= \frac{2}{\pi} \int_{-\frac{\pi}{2}}^{\frac{\pi}{2}} \frac{1}{4} \sin^2(2\varphi) \, d\varphi \\
 &\stackrel{\psi=2\varphi}{=} \frac{1}{2\pi} \int_{-\pi}^{\pi} \sin^2(\psi) \frac{1}{2} \, d\psi \\
 &= \frac{1}{4\pi} \int_{-\pi}^{\pi} \frac{1}{2} (\sin^2(\psi) + \cos^2(\psi)) \, d\psi \\
 &= \frac{1}{8\pi} 2\pi \\
 &= \frac{1}{4}
 \end{aligned}$$

Damit folgt

$$\mathbb{V}\text{ar}(X) = \frac{1}{4}$$

Ganz analog

$$\text{Var}(Y) = \frac{1}{4}$$

Betrachte nun zunächst

$$\int_0^{2\pi} \sin(\varphi) \cos(\varphi) \, d\varphi \stackrel{\text{part. Integrat.}}{=} \underbrace{\sin^2 \varphi \Big|_0^{2\pi}}_{=0} - \int_0^{2\pi} \sin \varphi \cos \varphi \, d\varphi$$

Damit folgt

$$\int_0^{2\pi} \sin \varphi \cos \varphi \, d\varphi = 0 \quad (1)$$

Es ist $\int_{\mathbb{R}^2} \left| \frac{xy}{\pi} \right| \mathbb{1}_{\{(x,y) \in E\}} \, d(x,y) \leq \frac{1}{\pi} \mathcal{L}^2(E) = 1 < \infty$, d.h. Fubini ist anwendbar. Damit folgt

$$\begin{aligned} \mathbb{E}(XY) &= \int_{\mathbb{R}^2} xy f^{X,Y}(x,y) \, d(x,y) \\ &= \int_{\mathbb{R}} \int_{\mathbb{R}} \frac{xy}{\pi} \mathbb{1}_{\{(x,y) \in E\}} \, dx \, dy \\ &\stackrel{\text{Trafosatz}}{=} \frac{1}{\pi} \int_0^1 dr \int_0^{2\pi} d\varphi r^3 \cos(\varphi) \sin(\varphi) \\ &= \frac{1}{4\pi} \int_0^{2\pi} \cos(\varphi) \sin(\varphi) \, d\varphi \\ &\stackrel{(1)}{=} 0 \end{aligned}$$

Da $\mathbb{E}(X) = \mathbb{E}(Y) = 0$ folgt also

$$\text{Cov}(X, Y) = 0$$

Und damit

$$\rho(X, Y) = 0.$$

- (c) Es gilt nach VL: $X \perp\!\!\!\perp Y \iff f^{X,Y}(x,y) = f^X(x)f^Y(y)$ $\mathcal{L}$ -f.ü. Nun betrachte $A := (-1, 1)^2 \setminus E$. Dann ist

$$\mathcal{L}^2(A) = \mathcal{L}^2(A) - \mathcal{L}^2(E) = 2^2 - \pi = 4 - \pi > 0.$$

Also ist A keine $\mathcal{L}$ -Nullmenge. Jedoch gilt $\forall (x,y) \in A$:

$$f^{X,Y}(x,y) = 0 \neq \frac{4}{\pi^2} \underbrace{\sqrt{1-x^2}}_{>0} \underbrace{\sqrt{1-y^2}}_{>0}.$$

Also folgt X und Y nicht unabhängig.

Aufgabe 34. (a) Es gilt

$$\begin{aligned} \mathbb{F}^{Z_p}(x) &= \mathbb{P}^{Z_p}((-\infty, x]) \\ &= \mathbb{P}^{(-1)^{V_p} \cdot Y}((-\infty, x] \cap \{V_p = 0\}) + \mathbb{P}^{(-1)^{V_p} \cdot Y}((-\infty, x] \cap \{V_p = 1\}) \\ &= \mathbb{P}^Y((-\infty, x] \cap \{V_p = 0\}) + \mathbb{P}^{-Y}((-\infty, x] \cap \{V_p = 1\}) \\ &\stackrel{Y \perp\!\!\!\perp V_p}{=} \mathbb{P}(\{Y \leq x\}) \cdot \mathbb{P}(\{V_p = 0\}) + \mathbb{P}(\{-Y \leq x\}) \cdot \mathbb{P}(\{V_p = 1\}) \\ &\stackrel{\text{Symmetrie } N_{(0,1)}}{=} \mathbb{P}(\{Y \leq x\}) \cdot \mathbb{P}(\{V_p = 0\}) + \mathbb{P}(\{Y \leq x\}) \cdot \mathbb{P}(\{V_p = 1\}) \\ &= \mathbb{P}(\{Y \leq x\}) \cdot \mathbb{P}(\{V_p = 0\} \cup \{V_p = 1\}) \\ &= \mathbb{P}(\{Y \leq x\}) \\ &= \mathbb{F}^Y(x) \end{aligned}$$

Daher gilt $Z_p \sim Y \sim N_{(0,1)}$.

(b) Es gilt

$$\begin{aligned}\mathbb{P}(\{Y < -1, Z_p < -1\}) &= \mathbb{P}(\{Y < -1\} \cap \{V_p = 0\}) \stackrel{Y \perp\!\!\!\perp V_p}{=} \mathbb{P}(\{Y < -1\}) \cdot \mathbb{P}(\{V_p = 0\}) \\ &= (1-p) \cdot \mathbb{P}(\{Y < -1\})\end{aligned}$$

und völlig analog

$$\begin{aligned}\mathbb{P}(\{Y < -1, Z_p > 1\}) &= \mathbb{P}(\{Y < -1\} \cap \{V_p = 1\}) \stackrel{Y \perp\!\!\!\perp V_p}{=} \mathbb{P}(\{Y < -1\}) \cdot \mathbb{P}(\{V_p = 1\}) \\ &= p \cdot \mathbb{P}(\{Y < -1\})\end{aligned}$$

Angenommen, $Y \perp\!\!\!\perp Z_p$. Dann gilt

$$\begin{aligned}\mathbb{P}(\{Y < -1, Z_p < -1\}) &= \mathbb{P}(\{Y < -1\})\mathbb{P}(\{Z_p < -1\}) \\ (1-p) \cdot \mathbb{P}(\{Y < -1\}) &\stackrel{(a)}{=} \mathbb{P}(\{Y < -1\})^2 \\ (1-p) &= \mathbb{P}(\{Y < -1\})\end{aligned}$$

und völlig analog

$$\begin{aligned}\mathbb{P}(\{Y < -1, Z_p > 1\}) &= \mathbb{P}(\{Y < -1\})\mathbb{P}(\{Z_p > 1\}) \\ p \cdot \mathbb{P}(\{Y < -1\}) &\stackrel{(a), \text{Symmetrie } N_{(0,1)}}{=} \mathbb{P}(\{Y < -1\})^2 \\ p &= \mathbb{P}(\{Y < -1\})\end{aligned}$$

Nun führen wir eine Fallunterscheidung durch. Für $p = \frac{1}{2}$ folgt $\mathbb{P}(\{Y < -1\}) < \mathbb{P}(\{Y < 0\}) \leq \frac{1}{2}$, Widerspruch zu $p = \mathbb{P}(\{Y < -1\})$. Für $p \neq \frac{1}{2}$ erhalten wir aus $p = \mathbb{P}(\{Y < -1\}) = (1-p)$ ebenfalls einen Widerspruch. Daher ist $Y \not\perp\!\!\!\perp Z_p$.

(c) Es gilt

$$\begin{aligned}\mathbb{E}(YZ_p) &= \int_{\mathbb{R}} \int_{\mathbb{R}} y z \mathbb{f}^{Y, Z_p}(y, z) \, dy \, dz \\ &\stackrel{Z_p = (-1)^{V_p} Y}{=} \int_{0,1} \int_{\mathbb{R}} y^2 (-1)^v \mathbb{f}^{Y, V_p}(y, v) \, dy \, dv \\ &\stackrel{Y \perp\!\!\!\perp V_p}{=} (1-p) \cdot \int_{\mathbb{R}} y^2 \mathbb{f}^Y(y) \, dy + p \cdot \int_{\mathbb{R}} -y^2 \mathbb{f}^Y(y) \, dy \\ &= (1-2p) \int_{\mathbb{R}} y^2 \mathbb{f}^Y(y) \, dy\end{aligned}$$

Außerdem gilt

$$\begin{aligned}\mathbb{E}(y)\mathbb{E}(Z_p) &= \int_{\mathbb{R}} y \mathbb{f}^Y(y) \, dy \int_{\mathbb{R}} z \mathbb{f}^{Z_p}(z) \, dz \\ &\stackrel{Z_p = (-1)^{V_p} Y}{=} \int_{\mathbb{R}} y \mathbb{f}^Y(y) \, dy \cdot \int_{0,1} \int_{\mathbb{R}} y (-1)^v \mathbb{f}^{Y, V_p}(y, v) \, dy \, dv \\ &\stackrel{Y \perp\!\!\!\perp V_p}{=} (1-p) \cdot \left(\int_{\mathbb{R}} y \mathbb{f}^Y(y) \, dy \right)^2 - p \cdot \left(\int_{\mathbb{R}} y^2 \mathbb{f}^Y(y) \, dy \right)^2 \\ &= (1-2p) \left(\int_{\mathbb{R}} y \mathbb{f}^Y(y) \, dy \right)^2\end{aligned}$$

Für $p = 0$ erhalten wir daher

$$\text{Cov}(Y, Z_p) = \mathbb{E}(YZ_p) - \mathbb{E}(Y)\mathbb{E}(Z_p) = 0 - 0 = 0.$$

Aufgabe 35. (a) Rechnen ergibt für $z \in \mathbb{R}$

$$\begin{aligned}
 \mathbb{F}^{M_1}(z) &= \mathbb{P}(\{M_1 \leq z\}) \\
 &= \mathbb{P}\left(\left\{\min_{i \in \{1, \dots, n\}} X_i \leq z\right\}\right) \\
 &= 1 - \mathbb{P}\left(\bigcap_{i=1}^n \{X_i > z\}\right) \\
 &\stackrel{\text{unabh.}}{=} 1 - \prod_{i=1}^n \mathbb{P}(\{X_i > z\}) \\
 &\stackrel{\text{idv}}{=} 1 - (1 - \mathbb{F}^X(z))^n \\
 \mathbb{F}^{M_2}(z) &= \mathbb{P}(\{M_2 \leq z\}) \\
 &= \mathbb{P}\left(\prod_{i=1}^n \{X_i \leq z\}\right) \\
 &\stackrel{\text{unabh.}}{=} \prod_{i=1}^n \mathbb{P}(\{X_i \leq z\}) \\
 &\stackrel{\text{idv}}{=} \mathbb{F}^X(z)^n.
 \end{aligned}$$

(b) Für $X_1 \sim \text{Exp}_\lambda$ ist $\mathbb{F}^{X_1} = (1 - \exp(-\lambda z)) \mathbb{1}_{\mathbb{R}^+}$. Damit folgt

$$\begin{aligned}
 \mathbb{F}^{M_1}(z) &= 1 - (1 - \mathbb{F}_{\text{Exp}_\lambda}(z))^n \\
 &= 1 - [1 - (1 - \exp(-\lambda z)) \mathbb{1}_{\mathbb{R}^+}(z)]^n \\
 &= 1 - [1 - \mathbb{1}_{\mathbb{R}^+}(z) + \exp(-\lambda z) \mathbb{1}_{\mathbb{R}^+}(z)]^n \\
 &= 1 - \begin{cases} \exp(-\lambda n z) & z \in \mathbb{R}^+ \\ 1 & z \notin \mathbb{R}^+ \end{cases} \\
 &= (1 - \exp(-\lambda n z)) \mathbb{1}_{\mathbb{R}^+} \\
 &= \mathbb{F}_{\text{Exp}_{n\lambda}}.
 \end{aligned}$$

Da die Verteilungsfunktion das W-Maß eindeutig festlegt, folgt $M_1 \sim \text{Exp}_{n\lambda}$.

(c) Für $X_1 \sim U_{[0, \theta]}$ ist die Dichte $f(x) = \frac{1}{\theta} \mathbb{1}_{[0, \theta]}$ gegeben. Damit folgt

$$\begin{aligned}
 \mathbb{E}_\theta(X_1) &= \int_{\mathbb{R}} \frac{x}{\theta} \mathbb{1}_{[0, \theta]} dx \\
 &= \frac{1}{\theta} \int_0^\theta x dx \\
 &= \frac{\theta}{2} \\
 \mathbb{E}_\theta(X_1^2) &= \int_{\mathbb{R}} \frac{x^2}{\theta} \mathbb{1}_{[0, \theta]} dx \\
 &= \frac{1}{\theta} \int_0^\theta x^2 dx \\
 &= \frac{\theta^2}{3} \\
 \text{Var}_\theta(X_1) &= \mathbb{E}_\theta(X_1^2) - \mathbb{E}_\theta(X_1)^2 \\
 &= \frac{\theta^2}{3} - \frac{\theta^2}{4} \\
 &= \frac{\theta^2}{12}.
 \end{aligned}$$

Es gilt $f^{M_2} = (\mathbb{F}^{M_2})'$. Damit folgt für $z \in \mathbb{R}$:

$$\begin{aligned} f^{M_2}(z) &= (\mathbb{F}^{M_2})'(z) \\ &= (\mathbb{F}^X(z)^n)' \\ &= n(\mathbb{F}^X(z)^{n-1})f^X(z) \\ &= n \left(\frac{(z \wedge \theta) \vee \theta}{\theta} \right)^{n-1} \frac{1}{\theta} \mathbb{1}_{[0, \theta]}(z) \\ &= \frac{n}{\theta^n} z^{n-1} \mathbb{1}_{[0, \theta]}(z) \end{aligned}$$

Damit folgt

$$\begin{aligned} \mathbb{E}_\theta(M_2) &= \int_{\mathbb{R}} x \frac{nx^{n-1}}{\theta^n} \mathbb{1}_{[0, \theta]} dx \\ &= \frac{n}{\theta^n} \int_0^\theta x^n dx \\ &= \frac{n}{n+1} \theta. \end{aligned}$$

(d) Rechnen ergibt

$$\begin{aligned} \mathbb{E}_\theta(\overline{X}_n) &= \mathbb{E}_\theta \left(\frac{1}{n} \sum_{k=1}^n X_k \right) \\ &\stackrel{\text{lin.}}{=} \frac{1}{n} \sum_{k=1}^n \mathbb{E}_\theta(X_k) \\ &\stackrel{\text{idv}}{=} \mathbb{E}_\theta(X_1) \\ &= \frac{\theta}{2} \end{aligned}$$

Damit folgt

$$\begin{aligned} \text{Bias}_\theta(\hat{\theta}_1) &= \mathbb{E}_\theta(\hat{\theta}_1 - \theta) \\ &= 2\mathbb{E}_\theta(\overline{X}_n) - \theta \\ &= 2\frac{\theta}{2} - \theta \\ &= 0 \\ \text{Bias}_\theta(\hat{\theta}_2) &= \mathbb{E}_\theta(\hat{\theta}_2 - \theta) \\ &= \mathbb{E}_\theta(M_2) - \theta \\ &= \frac{n}{n+1} \theta - \theta \\ &= -\frac{\theta}{n+1} \end{aligned}$$

Nun rechne

$$\begin{aligned} \mathbb{E}_\theta(M_2^2) &= \int_{\mathbb{R}} x^2 \frac{nx^{n-1}}{\theta^n} \mathbb{1}_{[0, \theta]}(x) dx \\ &= \frac{n}{\theta^n} \int_0^\theta x^{n+1} dx \\ &= \frac{n}{n+2} \theta^2 \end{aligned} \tag{2}$$

Es ist offensichtlich $\hat{\theta}_3$ nun erwartungstreu. Damit folgt für $n > 1$

$$\begin{aligned}
 \text{Var}_\theta(\hat{\theta}_1) &= 4\text{Var}_\theta(\bar{X}_n) \\
 &\stackrel{\text{unabh.}}{=} \frac{4}{n^2} \sum_{k=1}^n \text{Var}_\theta(X_k) \\
 &\stackrel{\text{idv}}{=} \frac{4}{n} \text{Var}_\theta(X_1) \\
 &= \frac{\theta^2}{3n} \\
 &> \frac{\theta^2}{n(n+2)} \\
 &= \left(\frac{n+1}{n} \right)^2 \frac{n}{n+2} \theta^2 - \theta^2 \\
 &\stackrel{(2)}{=} \left(\frac{n+1}{n} \right)^2 \mathbb{E}_\theta(M_2^2) - \mathbb{E}_\theta \left(\frac{n+1}{n} M_2 \right)^2 \\
 &= \mathbb{E}_\theta(\hat{\theta}_3^2) - \mathbb{E}_\theta(\hat{\theta}_3)^2 \\
 &= \text{Var}_\theta(\hat{\theta}_3)
 \end{aligned}$$

Schlussendlich ergibt sich

$$\begin{aligned}
 \text{MSE}_\theta(\hat{\theta}_1) &= \mathbb{E}_\theta(|\hat{\theta}_1 - \theta|^2) \\
 &= 4\mathbb{E}_\theta(\bar{X}_n^2) - \theta^2 \\
 &= 4\text{Var}_\theta(\bar{X}_n) + 4\mathbb{E}_\theta(\bar{X}_n)^2 - \theta^2 \\
 &\stackrel{\text{iid}}{=} \frac{4}{n} \text{Var}_\theta(X_1) \\
 &= \frac{\theta^2}{3n} \\
 \text{MSE}_\theta(\hat{\theta}_2) &= \mathbb{E}_\theta(\hat{\theta}_2^2) - 2\theta\mathbb{E}_\theta(M_2) + \theta^2 \\
 &= \frac{n}{n+2} \theta^2 - 2 \frac{n}{n+1} \theta^2 + \theta^2 \\
 &= \frac{2\theta^2}{(n+2)(n+1)} \\
 \text{MSE}_\theta(\hat{\theta}_3) &= \mathbb{E}_\theta(\hat{\theta}_3^2) - 2\theta\mathbb{E}_\theta(\hat{\theta}_3) + \theta^2 \\
 &= \left(\frac{n+1}{n} \right)^2 \frac{n}{n+2} \theta^2 - \theta^2 \\
 &= \frac{\theta^2}{n(n+2)}.
 \end{aligned}$$